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Abstract

This paper re-examines the role of labor-market competition as a determinant of attitudes toward immigration. We claim two main contributions. First, we use more sophisticated measures of the degree of exposure to competition from immigrants than previously done. In addition to education, we focus on the protection derived from investments in job-specific human capital and from specialization in communication-intensive tasks. Second, we explicitly account for the potential endogeneity arising from job search. Methodologically, we estimate, by instrumental variables, an econometric model that allows for heterogeneity at the individual, regional, and country level. Drawing on the 2004 European Social Survey, we obtain the following main results: First, we find that individuals with above-average dislike for immigrants tend to work in low-immigration jobs, biasing OLS estimates downwardly. Second, we show that individuals who are currently employed in jobs that require high levels of specific human capital and/or high communicational skills are relatively more pro-immigration. Third, we find that the protection granted by job-specific human capital is clearly distinct from the protection granted by formal education. In contrast, the positive effect of education on pro-immigration attitudes is greatly reduced when we control for the degree of communication intensity of respondents' occupations. Finally, we find that the effect of communication (manual) intensity on immigration preferences operates at different levels of education. Overall our results suggest a large role for skill-based labor market competition in determining individual attitudes toward immigration.

Key Words: Immigration Attitudes, Labor Market, Job-Specific Human Capital, Communication Skills, International Migration. JEL Codes: F1, F22, J61, J31, R13.

INTRODUCTION*

Over the last decade a great deal of research has been devoted to analyzing the economic determinants of attitudes toward immigration. Special attention has been paid to testing the hypothesis that competition with immigrants in the labor market influences native workers' views.¹

The degree of competition between native and immigrant workers is defined by their relative skills. In most studies skills are measured using education levels. Under the assumption that immigrants are, on average, less educated than natives, it is expected that low-educated natives oppose immigration to avoid depressing their wages. In contrast, being less exposed to competition from immigrants, highly educated natives should hold less negative views toward (low-educated) immigration and could even be in favor of it, depending on the elasticity of the supply of capital. Typically, most studies have found a positive association between respondents' educational attainment and their pro-immigration views. This association has been interpreted as evidence that labor-market competition has indeed attitudinal effects.

We find such interpretation wanting for the following two reasons. First, it must be noted that the observed relationship between education and attitudes toward immigration does not constitute by itself unequivocal evidence in favor of the labor market exposure hypothesis since mechanisms other than competition could

be driving this association. Sociologists and political scientists have long been putting forward alternative interpretations of the attitudinal effects of education (see e.g. Jackman and Muha 1981, Bobo and Licari 1989, Burns and Gimpel 2000, Kingston *et al.* 2003). They have argued that the positive effect of education on pro-immigration attitudes could be entirely due to values and predispositions that are associated to schooling, such as tolerance, open-mindedness or political correctness, rather than to relative scarcity of human capital (see also Hainmueller and Hiscox 2007, Côté and Erickson 2009).

Secondly, and most importantly, we believe that defining skills solely in terms of years of education constitutes a very narrow definition of the human-capital resources that characterize native-immigrant competition in the labor market and hence provides an incomplete test for the labor-market exposure hypothesis.

The main goal of our paper is to provide a re-examination of the role of labor-market exposure to competition as a determinant of attitudes toward immigration. The key feature in our analysis is that we consider a more comprehensive definition of skills than previously used in the literature. Whilst most studies define skills using educational levels only —typically distinguishing between college-graduates and workers without tertiary education—, we consider two further measures of skills influencing the degree of exposure to labor-market competition. These new measures are directly linked to the characteristics of the tasks workers perform at their jobs. From an empirical point of view, an attractive feature of these measures is their large variation, across jobs and occupations.

Our first new measure of labor-market exposure is a self-assessed measure of job-specific human capital. Workers acquire job-specific skills at their firms via formal training, informal instruction or learning-by-doing. Investing in job-specific skills makes workers less replaceable, regardless of their level of education (Becker 1993[1964], Lazear 1995). Job-specific human capital thus provides protection against competition from other workers, native and immigrant alike.

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¹ Mayda (2006) provides a useful brief overview of the relevant literature.

Our second measure of labor market exposure is based on objective characteristics of respondents' occupations and is motivated by recent work studying how native workers respond to recent immigration. Peri and Sparber (2009) show that when immigrants arrive into an economy, native workers mitigate the wage effects of immigration by shifting toward occupations for which they have a comparative advantage. Specifically, immigration induces native workers to shun manual jobs and specialize in communication-intensive occupations. Natives' comparative advantage in such occupations stems from the possession of those skills that immigrants typically lack, in particular, language and/or cultural-specific skills. We hypothesize that native workers in communication-intensive (manual) occupations will be more (less) protected from immigrants' competition and hence will be more (less) likely to display pro-immigration attitudes.^{2,3}

By estimating *directly* the attitudinal effects of respondents' general, job-specific and communicational skills, we attempt to cover all the skill dimensions of labor-market competition. To our knowledge, ours is the first paper that provides such a comprehensive test for the exposure hypothesis.

Another attractive feature of our analysis is the use of individual-level data for many countries, as in Mayda (2006). Our main data source is the 2004 European Social Survey. The immigration module in this dataset has several unique features. It contains detailed questions that are highly relevant to understanding individual opinions on immigration. In addition, identical questions are posed to all countries included in the survey, allowing for high quality cross-country comparisons.⁴ For our

purposes, another attractive feature of the data is that it identifies the respondents' region of residence within the country. Following Autor, Levy and Murnane (2003) and Peri and Sparber (2009), we also use the US Occupational Network Online Dataset (O*NET) to build measures of the intensity of interactive (and manual) tasks by occupation.

Methodologically, we estimate an econometric model that allows for heterogeneity at the individual, regional, and country level. The dependent variable is a measure of the respondent's views toward immigration and we consider three dimensions of skills that determine these views: formal education, required job-learning time at the current job net of education (that is, job-specific human capital) and communication-intensity of the current occupation. Each of these skill variables measures a distinct source of protection from immigrant competition in the labor market. We expect lower exposure in each dimension to be associated with more favorable immigration attitudes.

Another important contribution of our study is our treatment of individual heterogeneity. Our analysis explicitly accounts for heterogeneity among natives in their views toward immigration and for potential self-selection into low-immigration jobs. Individual observable heterogeneity is addressed by estimating specifications that include a vector of controls for ideological and attitudinal variation, whilst self-selection is addressed by using an instrumental-variable approach. Specifically, we use the *regional* availability of low-exposure jobs as an instrument for actual *individual* exposure in the current job. To our knowledge, building an instrument for individual labor market exposure is novel in the literature.⁵

Overall, our results suggest a larger role for labor market competition as a

² Supply constraints or job searching costs may limit the number of native workers that can shift from manual to communication-intensive occupations.

³ Peri and Sparber (2009) use data for the US. Amuedo-Dorantes and De la Rica (2009) conduct a similar study using data for Spain and focusing on gender differences.

⁴ Card, Dustmann and Preston (2005) provide a detailed overview of the immigration module in the 2002 European Social Survey.

⁵ Endogeneity concerns of the type addressed here are also at the heart of the analysis of Dustmann and Preston (2001). They are interested in the effects of ethnic concentration at the local level on individual attitudes toward immigration. To that effect they build an instrument for local ethnic concentration using regional data.

determinant of individual attitudes toward immigration than previously found, as well as a new interpretation for the positive association between education and pro-immigration attitudes. More specifically, we report the following four main findings.

First, the limited role for labor market competition in earlier studies may have been due to a combination of poor skill indicators and endogeneity problems. Our estimates suggest that individuals with above-average dislike for immigrants tend to work in low-immigration jobs, biasing OLS estimates downwardly.

Second, our instrumental-variables estimates show that individuals employed in less exposed jobs are relatively more pro-immigration. This is true for our two new measures of exposure. Both the amount of specific human capital required by the job as well as the type of skills required in any given occupation play a crucial role in shaping attitudes toward immigration. This finding provides indirect evidence in support of the mechanism in Peri and Sparber (2009). Namely, it is consistent with the prediction that unskilled natives will shift away from jobs and occupations where they face greater competition from immigrant workers.

Third, the positive effect of formal education on pro-immigration attitudes is orthogonal to job-specific human capital but it is greatly reduced when we control for the level of communication-intensive (manual) tasks in the current occupation. This suggests that, to some extent, education is a proxy for the type of occupation.

Fourth, the positive effect of communication intensity on pro-immigrant preferences operates both at low and high levels of education. This indicates that communicational intensity is indeed a distinct source of labor-market protection vis-à-vis immigrant workers.

The paper is organized as follows. Section 2 situates this paper in the context of the previous literature. Section 3 presents the data sources, definitions of variables and descriptive statistics. Section 4 introduces our estimation method and explains how we deal with endogeneity. The main findings are presented in Section

5. Sensitivity analyses are reported in Section 6. Section 7 concludes.

PREVIOUS LITERATURE

Our paper is part of a large and growing body of literature analyzing the determinants of attitudes toward immigration. We can classify the work that is more directly related to our analysis in three main groups.

First, several authors have attempted to quantify the contribution of labor market considerations relative to welfare state considerations and to non-economic factors in explaining attitudes toward immigration. Using factor analysis, Dustmann and Preston (2005) analyze the determinants of immigration attitudes using the 2002 European Social Survey. With a similar methodology, Dustmann and Preston (2007) use data for Great Britain. To identify the role of labor market concerns they employ survey questions on fear of job loss, ease of finding a job, and expected future earnings. The results are similar in both cases. They find that subjective labor market concerns are a significant determinant of attitudes toward immigration. However, fiscal considerations and cultural and racial concerns seem to play a larger role. We note that their labor market variables are not directly related to exposure to competition from immigrants.

Secondly, our paper is closely related to the large empirical literature examining the relationship between individual education levels and attitudes toward immigration. The common finding across all papers listed below is that more educated individuals are more pro-immigration, a finding that is typically interpreted as evidence in favor of the labor-market competition hypothesis. Among the early studies, Espenshade and Hempstead (1996) and Scheve and Slaughter (2001) use data for the US, while Dustmann and Preston (2001, 2005) study the UK. More recently, Mayda (2006) and O'Rourke and Sinnott (2006) have employed individual-level data covering several (mostly rich) countries. Mayda (2006) shows that education is more strongly associated to pro-immigration attitudes in countries with higher GDP per

capita. She provides a labor-market interpretation for her finding, which is consistent with the factors proportion model. O'Rourke and Sinnott (2006) find a similar result. They also find that non-economic considerations play a larger role than labor market concerns. Facchini and Mayda (2009) study empirically the joint relationship between individual income and education and attitudes toward immigration. They find that in countries with relatively unskilled immigration, individual income is negatively correlated with pro-immigration preferences, while education has a positive effect. They propose a model with endogenous income redistribution that can rationalize this pattern.⁶

Finally, in addressing endogeneity in the analysis of attitudes toward immigration our paper connects to the work of Dustmann and Preston (2001) on the effects of ethnic concentration at the local level. These authors argue that the location choices of immigrants (and natives) are likely to depend on unobserved determinants of individual attitudes. They propose an instrumental variables approach, where *region*-level ethnic composition is used as an instrument for *local* ethnic composition. Controlling for endogeneity, they find that high local ethnic concentration leads to worse attitudes toward immigration among natives. Their results suggest that ignoring the endogeneity problem leads to underestimating the effect, since it appears that natives that dislike immigrants tend to locate in low-immigration localities.

DATA

Sources

Our main data source is the 2004 European Social Survey (ESS). It contains information on over 47,000 individuals from 25 countries. All countries with the exception of Luxemburg and Iceland are subdivided into regional units (273 regions). Most countries are 3-4% of the total sample each. Appendix 1 contains a

list of countries with the number of individual observations per country. Both the cross-country and the cross-region, within-country variation will be important in our analysis. We restrict our sample to currently employed individuals in age bracket 18-64 who are citizens in the respective country of residence in 2004. The resulting sample contains over 20,000 individuals.

We also use data from O*NET, which provides a detailed description of the characteristics of each occupation in the US. The O*NET dataset is developed under the sponsorship of the US Department of Labor/Employment and Training Administration. The O*NET dataset provides very detailed information of the mix of knowledge, skills and abilities required, as well as the activities and tasks typically performed, in 449 different occupations. O*NET provides 277 descriptors for each occupation. Each descriptor consists of a score ranging from 0 to 1 for each dimension of skill or ability considered. These scores are collected by occupation analysts and are constantly updated by ongoing surveys of each occupation's worker population and external occupation experts. We use O*NET skill descriptors to construct two measures of occupation-specific skills relating to labor-market competition between natives and immigrants. The first measure is an index of the amount of communicational (or interactive) skills required in each occupation, whilst the second measures the amount of physical (or manual) skills. These measures are subsequently matched to European occupational data under the assumption that occupations across US and European economies do not differ in relation to these two basic skill dimensions (see below).⁷

Definitions

In line with the existing literature, our dependent variable measures respondents' desired levels of immigration in their countries. The ESS contains three related questions. In each question, the respondent

⁶ Another theoretical model analyzing the determination of attitudes toward immigration in the presence of an endogenously determined welfare state is Ortega (2009).

⁷ For more details see <http://online.onetcenter.org/>.

is asked whether more immigrants of a particular origin should be allowed in the country. There are four possible responses: none, a few, some, or many. The three related questions differ on the origin of immigrants: the same race/ethnic group as the current majority in the population, a different race/ethnic group or from poorer countries outside Europe.

Our main dependent variable (IM1) is a simple transformation of the question referring to the desired level of immigration of individuals belonging to the *same* ethnic group as the current majority. We re-scale the variable so that it ranges from 0 to 100, where higher numbers mean higher support for immigration (i.e. lower entry restrictions). Our new variable takes 4 values: 25 (none), 50 (a few), 75 (some), and 100 (many).⁸ We note that this question can be interpreted as the respondent's views on how immigration affects his or her individual outcomes. By leaving ethnicity considerations aside, we believe that this variable is better equipped to capture the *economic* determinants of attitudes toward immigration. Yet we also analyze whether our main results are robust to variations in the definition of our dependent variable.

One of our main explanatory variables is a measure of the specific human capital required by the respondent's job, measured as job-learning time for individuals with the appropriate educational credentials. We note that this is a subjective measure as it is based on respondents' self-assessments, yet the very wording of the question is such that minimizes self-reporting bias. Specifically, the ESS asks respondents: "*If someone with the right education and qualifications replaced you in your job, how long would it take for them to learn to do the job reasonably well?*". We note that the question clearly differentiates between general and job-specific human capital. The original question allows for eight possible

answers, detailing intervals of time expressed in days, weeks, months or years. Based on this information we create a numerical variable expressed alternatively in days or months that can be used in a regression framework. We shall denote our variable by SHK (specific human capital). SHK has a mean of 265 days or, equivalently, 8.8 months. Interestingly, the correlation with education is positive but very low (0.13). This suggests that the wording of the question may indeed allow us to disentangle the effect of education from the effect of job-specific human capital on immigration views.

We also use a second set of variables to measure workers' degree of protection to labor market competition from immigrants. Following Peri and Sparber (2009), we assume that natives are better at communicational or interactive (relative to manual) tasks than recent immigrants. As explained-above, our measures of communication and manual skills were constructed using the O*NET dataset. We proceeded as follows: first, we used exploratory factor analysis to identify our skill dimensions of interest out of the 277 descriptors available in the dataset. Informed by this analysis, we constructed two different measures. Measure 1 is an index that tells us how communicational-intensive a particular occupation is. For each occupation of the dataset, this index simply averages the scores rating the task-importance and the mean observed abilities of 6 different communicational skills previously identified by factor analysis as part of the same skill dimension. These skills are: oral comprehension, oral expression, written comprehension, written expression, speech recognition, and speech clarity. Measure 2 captures, in turn, how important physical skills are in each occupation by averaging the task-importance and observed ability scores of the following 7 descriptors: visualization, arm-hand steadiness, manual dexterity, finger dexterity, control precision, wrist-finger speed and visual color discrimination.

These two indices are then matched to ISCO-88 occupations using the same occupational crosswalk employed by

⁸ Our scaling choice is arbitrary but our interpretation of the magnitudes of the estimated coefficients will be independent of this scale. We have also experimented with a dependent variable that combines several related questions on attitudes using factor analysis. Our main results were very similar. Those results are available upon request.

Amuedo-Dorantes and De la Rica (2009).⁹ Further details on the construction and merging of O*NET scores can be found in the Appendix. In the O*NET dataset the two indices (manual and communication intensities) are strongly negatively correlated, with a coefficient of correlation equal to -0.54. In the ESS data the correlation goes up to -0.78.¹⁰ This very high correlation implies that, effectively, both variables contain the same information: the degree of communication-intensive relative to manual tasks required by each occupation. Throughout the analysis we focus on the measure of communication intensity. In the robustness section we will also use the index of manual skills. Note that, implicitly, we are assuming that the skill requirements in US occupations are the same as the skill requirements in the same occupations in Europe. While it would certainly be a stretch to make this assumption for all the occupational descriptors included in O*NET or if we were comparing countries with widely different levels of economic development, we believe it is a reasonable assumption in the context of our analysis.

Descriptive Statistics

Let us now provide a brief description of the main variables in our analysis. We start with the dependent variable: a measure of the respondent's desired level of immigration of people of the same ethnicity as most of the population in the country (IM1). Appendix 2 contains the exact wording of the question. Appendix 3 reports some summary statistics. The modal score is 75, with 46% of the sample, and a total of 68% supporting some more or many more immigrants of the same ethnicity. As for the rest of the sample, 25% think that only a few more immigrants should be admitted and about 7% of the population thinks no more immigrants should be

allowed into their countries. The mean response across all countries (score 70.9) is very close to "allow some more immigrants." However, it is interesting to note that there is substantial variation across countries. Table 1a presents the mean values for the 25 countries. For Portugal, Turkey, Greece, and the Czech Republic the mean ranges from 56 to 64 (roughly, allow a few immigrants to enter). At the other end of the spectrum, Sweden, Ukraine and Iceland are the more pro-immigrant, with mean values in the 80-86 range. The second column in Table 1a reports attitudes toward immigration that are ethnically different to the majority of the population in the destination country. Clearly, pro-immigration attitudes are lower (average score 65) than for the case of immigrants with the same ethnicity as the majority of the population (average score 71). However, we note that the values in column 2 are approximately increasing as we move down the table, suggesting that the ranking of countries is practically the same for the two measures of attitudes toward immigration. We shall use this variable in our robustness section.

Let us now turn to the main explanatory variables. To relate to earlier literature it is helpful to begin by reviewing the data on years of education. As shown in column 3, the mean across all countries is 12.7 years, but the data display a larger variation, ranging from 8.7 (Turkey) to 14.4 (Denmark) years of education. Again we note that the values in column 3 are generally increasing as we move down the list, even though the relationship is not very strong. This illustrates that the positive association between pro-immigration attitudes and years of schooling is not only found across individuals within a country but also across countries.

Next, let us examine our measure of job-specific human capital. The mean value across all countries for how much time is needed for someone with the right qualifications to replace a worker 8.9 months (Table 1a). Again we observe substantial cross-country variation, ranging from 3.9 (Portugal) to 13.2 months (France), as illustrated by Table 1a. As we noted earlier, our measure of job-specific human capital is only weakly correlate with

⁹ The authors are most grateful to Sara De La Rica and Catalina Amuedo Dorantes for kindly making their crosswalk available to us.

¹⁰ The main reason for the increase in correlation is that the O*NET-ESS crosswalk only allows us to match 3-digit O*NET occupational codes into 2-digit ISCO-88 occupations.

TABLE 1a and 1b. Descriptives

TABLE 1a. Means by Country

Country	IM1	IM2	Education years	SHK in months	Comm. skills	Manual skills
Portugal	57.69	54.93	9.36	3.89	0.45	0.51
Greece	59.63	54.17	12.07	6.64	0.43	0.53
Turkey	59.98	52.01	8.73	13.04	0.39	0.55
Hungary	63.70	52.37	12.34	11.56	0.44	0.55
Czech Republic	63.78	59.95	12.87	6.35	0.43	0.56
Estonia	67.27	57.11	13.47	5.05	0.46	0.53
Great Britain	67.41	63.42	12.83	11.76	0.51	0.49
Finland	68.12	62.74	13.74	11.29	0.50	0.52
Netherlands	68.27	65.36	13.51	10.46	0.57	0.47
Spain	68.95	67.73	13.04	8.23	0.47	0.51
Slovenia	69.06	66.26	12.32	12.91	0.45	0.56
France	70.02	65.98	12.80	13.24	0.51	0.49
Austria	70.58	65.43	12.78	6.81	0.49	0.50
Luxembourg	70.64	63.78	13.15	9.99	0.51	0.49
Belgium	71.80	66.43	13.53	9.86	0.50	0.51
Germany	71.86	65.02	14.06	8.72	0.50	0.51
Poland	73.35	70.31	12.89	8.85	0.41	0.57
Slovakia	75.00	69.93	12.89	6.75	0.45	0.56
Norway	75.35	70.09	13.92	9.63	0.50	0.50
Ireland	75.47	71.18	13.40	7.91	0.48	0.51
Switzerland	76.94	71.60	10.93	7.06	0.51	0.51
Denmark	76.99	67.44	14.40	9.51	0.51	0.50
Sweden	80.78	79.08	13.06	8.79	0.49	0.52
Ukraine	85.09	72.86	13.05	7.13	0.48	0.53
Iceland	86.01	77.87	14.10	6.91	0.49	0.51
min	57.69	52.01	8.73	3.89	0.39	0.47
mean	70.95	65.32	12.77	8.89	0.48	0.52
median	70.58	65.98	13.04	8.79	0.49	0.51
max	86.01	79.08	14.40	13.24	0.57	0.57

Notes: Sorted from less to more pro-immigration based on the main dependent variable (IM1). SHK (specific-human capital) in months. See appendix for definitions of all variables.

years of education. The correlation coefficient is 0.13.

We now turn to the indices of the communication and manual skill requirements by occupation. By construction, these indices take values between zero and one. Across countries the mean value for the communication skills index ranges from 0.36 in Turkey or 0.38 in Poland to 0.54 in the Netherlands, with a mean value of 0.45. The manual skills index has a significantly shorter range, taking values between 0.48 (Netherlands) and 0.57-0.58 (Turkey and Poland). We also note that the top of the ranking by communication skills coincides with the bottom by manual skills, and vice versa. As we discuss later, the correlation coefficient between these two variables is very close to minus one.

Table 1b presents descriptive statistics for all the relevant variables for our main sample of individual observations. In the sample, immigration attitudes (IM1) range between 25 and 100, and years of education range from 0 to 25. Crucial to our identification strategy, there is high variation across individuals in job-specific human capital (SHK), and communication and manual skills. In terms of correlations, years of education is positively correlated with communication skills and negatively with manual skills. In turn, the communication skills index is positively correlated with job-specific human capital but this correlation is weak (correlation coefficient 0.15). Manual dexterity is not correlated with job-specific human capital (-0.02). The table also reports a number of individual-level variables that we will use to control for (observable) individual heterogeneity in immigration views. These variables are dummy variables for the presence of children in the house, for living in a rural area, for having a foreign-born mother, and a vector of individual attitudes and perceptions, including ideology, religiosity, happiness, trust, and social capital.

To provide some insight into what is captured by our measures of communication and manual skills, we next compare the samples of individuals currently employed in communication-intensive occupations (75 percentile) to

those in manual-intensive occupations (75 percentile). We find that, on average, manual workers have 2.6 fewer years of education, and 77 fewer days are required for someone with the right qualifications to do the job well.

Furthermore, Table 2 reports the top 10 occupations ranked by their average years of schooling, average job-specific human capital (job-learning time), average communicational intensity scores and average manual intensity scores. Two features are worth noting. First, several of the top occupations by educational attainment are also in the top ranking by communicational skills. For instance, medical doctors and biologists (life and health professionals), mathematicians and physicists, or higher-education teaching professionals appear in both lists. Second, the overlap is much smaller between the top occupations by educational attainment and by job-specific human capital. In fact the latter contain several occupations that do not require much formal education, such as “riggers and cable splicers”, “wood processing and paper-making plant operators”, or “upholsterers”. This observation suggests that specific and general human capital (in the form of years of education) are indeed two separate dimensions of labor market exposure.

ESTIMATION

Econometric Model

We estimate linear regression models.¹¹ Our dependent variable is a measure of individual views on immigration. Specifically, $IM(i,r,c)$ is the response of individual i living in region r and country c to the question of whether immigration should increase or decrease. Our model attempts to explain individual variation in this variable employing several models of the form:

¹¹ Using linear regression models simplifies the assessment of our instrumental variables strategy. Yet our results are robust to using discrete-choice models as in Mayda (2006) or O'Rourke, K. H. and R. Sinnott (2006). Estimates using discrete-choice models are shown on Appendix 5.

TABLE 1a and 1b. Descriptives

TABLE 1b. Summary Main Variables

Variable	Obs	Mean	Std. Dev.	Min	Max
IM1	20121	70.87	21.23	25	100
yearsedu	20619	12.84	3.58	0	25
shk_months	16747	8.84	13.90	0.02	73
shk_days	16747	265.12	417.14	0.50	2190
comm	20123	0.48	0.22	0.09	0.95
manual	20123	0.52	0.17	0.24	0.86
country	20619	12.16	7.23	1	25
region	20618	138.95	75.80	1	273
age	20619	41.64	11.12	18	64
female	20596	1.46	0.50	1	2
d_children	20619	0.53	0.50	0	1
d_rural	20619	0.37	0.48	0	1
d_motherfb	20619	0.07	0.26	0	1
ideology	20619	5.11	1.97	0	10
religiosity	20467	4.39	2.87	0	10
happy	20532	7.41	1.80	0	10
trust	20609	5.20	1.96	0	10
social capital	20619	4.92	1.48	1	7

Source: ESS 2004.

TABLE 2. The Top 10 Occupations According to Average Years of Education, Average Job-Specific Human Capital (Job-Learning Time), Average Communicational Intensity Scores and Average Manual Intensity Scores

Education (in years)	<i>Mean</i>	Job-specific human capital (in months)	<i>Mean</i>
Physicists, Chemists and related prof.	20.8	Riggers and cable splicers	73
Medical Doctors	19.4	Photographic-products machine operators	42.6
Biologists, botanists, zoologists & related	18.7	Farming and forestry advisers-technicians	39.6
Higher education teaching professionals	18.4	Tobacco preparers and tobacco products makers	38.6
Judges	17.8	Aircraft engine mechanics and fitters	38.4
Veterinarians	17.6	Wood processing and paper-making plant operators	36.6
Lawyers	17.6	Physical, mathematical and engineering science professionals	34.5
Dentists	17.4	General managers of small enterprises	31.9
Psychologists	17.4	Upholsterers and related workers	27.6
Mathematicians and related professionals	17.1	Production and operations managers in agriculture, hunting, forestry and fishing	26.5
Communicational skills	<i>Score</i>	Manual skills	<i>Score</i>
Legislators and senior officials	0,950	Precision handicraft & printing workers	0,860
Corporate managers	0,807	Other craft workers	0,841
Life science & health professionals	0,760	Metal machinery workers	0,830
Other professionals	0,759	Machine operators and assemblers	0,793
Teaching professionals	0,715	Stationary plant operators	0,756
General managers	0,711	Extraction and building trade workers	0,741
Physical, math.& engineering sci. prof.	0,687	Agriculture & fishery unskilled occupations	0,708
Other associate professionals	0,657	Physical & engineering associate profess.	0,693
Life science & health assoc. profess.	0,638	Drivers and mobile plant operators	0,693
Teaching associate professionals	0,560	Skilled agricultural and fishery workers	0,679

Notes: Average years of education in occupation and average job-learning time in occupations calculated using ISCO-88 4-digit coding. Average communication intensity and manual intensity scores using ISCO-88 2-digit coding.

Source: ESS 2004 and O*NET.

$$IM_{i,r,c} = \alpha_c + \beta_1 SHK_{i,r,c} + \beta_2 Comm_{i,r,c} + \gamma edu_{i,r,c} + \delta_{i,r,c} + \varepsilon_{i,r,c} \quad (1)$$

where the right-hand side contains country-specific intercepts, our new measures of exposure (job-specific human capital and communication skill requirements in the current occupation), years of education, and a set of controls.¹² The vector of controls always includes, age, age squared, gender, and dummies for the presence of children in the household, living in a rural area, and having a foreign-born mother. Occasionally, we will also introduce a set of additional variables aimed at capturing individual heterogeneity that may be relevant to understand the respondent's views on immigration (ideology, religiosity, happiness, trust, and social capital).¹³ For short we shall refer to this latter set of variables as individual attitudes. Finally, we allow the error term to be correlated across individuals living in the same region.

As argued above, together with general skills, job-specific human capital and communicational skills should provide protection against labor-market competition from immigrants. Hence we expect the coefficient on years of education (γ) and on our two additional measures of labor market protection to immigration (β_1 and β_2) to be positive.

Endogeneity

We are concerned with the following endogeneity problem. Individuals that particularly dislike immigration will search more intensively for jobs with few immigrants. These jobs will tend to display a high degree of protection from competition from immigrants. As a result, OLS estimates of our coefficients of interest (β_1 and β_2) are likely to be downwardly biased.

¹² In principle it would be possible to include region-specific intercepts. We shall do so in one of our OLS specifications. However, given the definition of our instrument, only country-specific intercepts can be included in our instrumental-variables specifications.

¹³ We are aware that some of these attitudinal variables may be endogenous. To evaluate this concern we will estimate all our main specifications with and without the vector of attitudes.

Our strategy to deal with this endogeneity problem is to use instrumental variables. Specifically, we postulate that individuals living in a region where there is a high availability of highly protected jobs are more likely to end up in one of these jobs than a comparable individual in a region where low-exposure jobs are scarce. Thus we propose to use a measure of the *regional* availability of protected jobs as an instrument for the degree of protection in an *individual's* job. In particular, we instrument an individual's level of job protection by the mean protection of the average worker in his or her region of residence. Our identifying assumption is that natives do not sort into regions (or countries) based on their views on immigration.¹⁴ Similarly, and in line with the literature, we also assume that educational attainment is not determined by one's views over immigration.

Let us now examine the relevance of our instrument. Table 3 reports the results of the first-stage regressions. We report several specifications, varying in the dependent variable. In column 1, the dependent variable is SHK, our measure of job-specific human capital. The right-hand side of the regression contains country-specific dummies, a series of controls (not shown in the table) that include years of education, and the average value of SHK in the respondent's region of origin. Standard errors are heteroskedasticity-robust and have been clustered at the region level. The point estimated of the variable of interest is 0.92, very precisely estimated and with a t-statistic of 47.99. In words, individuals residing in a region with a large availability of jobs requiring high specific human capital are more likely to end up in such jobs, provided that they have the right education level.

Columns 2 and 3 report analogous regressions where the dependent variables are, in turn, our measures of communication skills and manual skills. As expected, the main coefficient is positive and highly significant in both cases. In addition, the

¹⁴ There is little evidence of displacement of natives in response to immigration in US data (Card and DiNardo 2000, Card 2001) or in European data (Gonzalez and Ortega, 2009).

TABLE 3. First-Stage Regression

Column	(1)	(2)	(3)	(4)	(5)	(6)
Dep. var.	SHK	COMM.	MANUAL	SKH	COMM.	MANUAL
avSHK_r	0.92***			0.92***	-0.0014**	0.0012**
<i>tstat</i>	47.99			41.78	-2.44	2.73
avCOMM_r		0.56***		-7.792***	0.58***	
<i>tstat</i>		19.38		-4.75	19.55	
avMANUAL_r			0.69***			0.68***
<i>tstat</i>			28.07			23.81
Cragg-Donald F	2303.04	375.58	787.92	895.79	191.38	283.37
Observations	16369	19632	19632	16029	16029	16029

Notes:

- Standard errors clustered by region and robust to heteroskedasticity.

*** p<0.01, ** p<0.05, * p<0.1.

- All specifications include country-specific dummies, years of education, controls for gender, age, age squared, as well as dummies for children, rural, and foreign-born mother.

values of the F-statistics imply that we can clearly reject the null hypothesis of weak instruments. Columns 4 through 6 report the first-stage regressions that correspond to our specifications including two endogenous measures of exposure at a time. Again, the results suggest that our instruments are strong. We also experimented using only regions for which we have at least 25 individual observations and the predictive power of our instrument was largely unchanged.

MAIN RESULTS

Job-Specific Human Capital

We are interested in testing the following hypothesis. Individuals employed in jobs characterized by high requirements of specific human capital hold more favorable views on immigration. This is because job-specific skills protect insider native workers from outside competition. To test this hypothesis we estimate the regression model described in equation (1). The dependent variable is a measure of the respondent's pro-immigration views (IM1). The right-hand side of the model contains country (or region) fixed effects, the degree of job-specific human capital in the worker's current job, years of education,

and a number of controls. We cluster standard errors at the regional level.

Table 4 presents the OLS (columns 1-4) and IV estimates (columns 5 and 6). As a benchmark, column 1 reports the OLS estimate of a regression where attitudes toward immigration are solely a function of years of education and individual controls. As expected, the education coefficient is positive and quite large (1.06). We also note the negative coefficients on the female, children, and rural dummies, and the large and positive effect of having a foreign-born mother. The following columns include the main explanatory variable, job-specific human capital. The OLS estimate is small and we cannot reject a value of zero (columns 2-3). Column 4 includes region fixed effects, which hardly affect the estimates obtained in column 3.

Column 5 presents our preferred specification for the IV estimation. We find that individuals employed in jobs characterized by high job-specific human capital are more pro-immigration. The effect is quite large and highly significant (p-value 0.053). An increase in four months in the time required to replace a worker (conditional on education) has roughly the same effect on attitudes toward immigration as one additional year of education. In this

TABLE 4. Job-Specific Human Capital (SHK)

Column	(1)	(2)	(3)	(4)	(5)	(6)
Estimation	OLS	OLS	OLS	OLS	IV	IV
yearsedu	1.066*** [0.0572]	1.072*** [0.0612]	0.933*** [0.0599]	1.086*** [0.0579]	0.943*** [0.0911]	0.808*** [0.0910]
SHK		0.0158 [0.0119]	0.0141 [0.0117]	0.0107 [0.0119]	0.224* [0.116]	0.221* [0.116]
female	-0.830*** [0.316]	-0.555 [0.351]	-0.955*** [0.345]	-0.637* [0.353]	0.372 [0.618]	-0.0616 [0.595]
age	-0.00289 [0.108]	-0.0700 [0.116]	-0.0493 [0.113]	-0.0593 [0.117]	-0.132 [0.121]	-0.118 [0.120]
age2	0.000204 [0.00129]	0.00101 [0.00140]	0.000637 [0.00138]	0.000870 [0.00140]	0.00141 [0.00142]	0.00110 [0.00141]
d_children	-0.0214 [0.306]	-0.106 [0.327]	-0.156 [0.327]	-0.0847 [0.332]	-0.280 [0.346]	-0.327 [0.346]
d_rural	-2.168*** [0.333]	-2.041*** [0.354]	-2.236*** [0.363]	-1.585*** [0.369]	-1.967*** [0.353]	-2.154*** [0.362]
d_motherfb	2.947*** [0.686]	2.663*** [0.728]	2.818*** [0.733]	2.227*** [0.713]	2.618*** [0.712]	2.751*** [0.716]
Observations	20098	16369	16232	16369	16369	16232
R-squared	0.132	0.126	0.151	0.159	0.108	0.134
Fixed effects	country	country	country	region	country	country
Attitudes	no	no	yes	no	no	yes

Notes:

- Standard errors clustered by region and robust to heteroskedasticity.
- Standard errors in square brackets. *** p<0.01, ** p<0.05, * p<0.1.
- Dependent variable is IM1, the answer to the question “Do you think there should be no more immigration, a few more immigrants, some more immigrants, or many more immigrants?” Question refers to immigrants belonging to the current majority group.
- Measures of Individual Values: Ideology, Religiosity, Happiness, Trust, and social capital.

sense, the marginal effect is three times larger for specific-human capital than for years of schooling. It is worth noting that the effect we have uncovered is distinct from the effect of educational attainment. The IV estimate of an additional year of education is 0.97, which is not statistically different from its OLS estimate. Finally, column 6 presents an additional IV specification including the vector of attitudinal variables. These additional controls reduce by roughly 10% the effect of education but do not affect our estimate of the effect of job-specific human capital.

We also note that our results suggest that OLS estimates of the effect of specific

human capital were downwardly biased, as we had expected. Our interpretation is that individuals who dislike immigrants search more intensively for jobs that are highly protected from immigration. As a result individuals sort into jobs that, given their qualifications, require higher job-specific human capital. The endogeneity bias that we uncover is similar to the one documented by Dustmann and Preston (2001). Their results suggest that individuals sort spatially according to their attitudes toward minorities.

Communicational Skills

Job-specific human capital provides protection toward all outsiders to the job, natives and immigrants alike. Here we try to be more specific and focus on a characteristic that only offers protection toward recent immigrants, but not toward natives or immigrants that are already well established and have fully assimilated. Building on Peri and Sparber (2009), we assume that natives have comparative advantage in communication-intensive tasks relative to recent immigrants. Hence, we hypothesize that individuals employed in communication-intensive occupations will be relatively more pro-immigration. Conversely, individuals employed in eminently manual occupations are expected to be relative less pro-immigration. We stress that our identification is based on comparisons across individuals within a country, holding education levels and other relevant demographic characteristics constant.

To analyze this question we estimate the model in equation (1) but we now use measures of the communication skill-requirements of respondents' current occupations. The results are presented in Table 5. Columns 1-4 report OLS estimates. Note that the estimate of years of education ranges between 0.61 and 0.74 after introducing our index of communication intensity. This is around 30% lower than in the previous section. In turn, the point estimate associated to communication skills is positive and significant, with values around 10. Columns 5 and 6 present our IV estimates. As was the case for job-specific human capital (Table 4), IV estimates of the measure of protection are substantially higher than OLS estimates, with a value of 27.06 in column 5. A simple calculation shows that an increase in the communication skills index of one standard deviation (0.22, Table 1b) leads to a 6-point increase in the immigration opinion index of 8.2. Column 6 shows that controlling for individual values and attitudes reduces further the effect of education but does not affect our estimate of the effect of communication skills. Finally, we also point out that years

of education is not significant in our IV estimates.

Education, Job-Specific Human Capital, and Communicational Skills

As discussed in the Introduction, most previous studies have focused on a single measure of human capital, namely, years of education. An important feature of our analysis is that we have considered several dimensions of skills. In the previous sections we have found that both job-specific human capital and communication skills are important determinants of individuals' attitudes toward immigration.

The goal of this section is to analyze the interaction between our two new measures of labor-market exposure and years of education. Table 6 presents the instrumental-variables estimates. As a benchmark, column 1 includes only years of education as explanatory variable (in addition to the demographic controls). The point estimate on years of education is 1.066. Including job-specific human capital (column 2) reduces the coefficient on education only slightly, suggesting that job-specific and general human capital (education) have differentiated effects on immigration views. The point estimate of SHK is 0.22 and the associated standard error is 0.12 (that is, a p-value of 0.053). As argued in the previous section, the effect of job-specific human capital is roughly three times larger than that of education, when both are measured in the same time units.¹⁵ Column 3 adds our measure of communication skills. The coefficient on job-specific human capital remains practically unchanged (0.21) but the coefficient on years of education drops dramatically (0.23) and becomes not statistically different from zero. Column 4 adds the vector of individual attitudes. The effect of education drops even further (0.09) whereas the coefficient on SHK remains practically unchanged (0.20). Finally, column 5 replaces communication skills with manual skills. The results largely confirm the message of the previous column. Naturally, the estimated effect of

¹⁵ Recall that SHK is measured in months whereas education is measured in years.

TABLE 5. Communication Intensity in Current Occupation

Column Estimation	(1) OLS	(2) OLS	(3) OLS	(4) OLS	(5) IV	(6) IV
yearsedu	1.066*** [0.0572]	0.714*** [0.0586]	0.607*** [0.0577]	0.744*** [0.0543]	0.218 [0.310]	0.114 [0.306]
comm.		11.18*** [0.908]	10.46*** [0.908]	10.68*** [0.889]	27.06*** [10.04]	26.46*** [10.07]
Observations	20098	19632	19431	19632	19632	19431
R-squared	0.132	0.141	0.165	0.170	0.122	0.145
Fixed effects	country	country	country	region	country	country
Attitudes	no	no	yes	no	no	yes

Notes:

- Standard errors clustered by region and robust to heteroskedasticity.
- Standard errors in square brackets. *** p<0.01, ** p<0.05, * p<0.1.
- Dependent variable is IM1, the answer to the question “Do you think there should be no more immigration, a few more immigrants, some more immigrants, or many more immigrants?” Question refers to immigrants belonging to the current majority group.
- All specifications include dummies for children, rural, and foreign-born mother. Also controls for gender, age, age squared, and dummies for children, rural and having a foreign-born mother.
- Measures of Attitudes: Ideology, Religiosity, Happiness, Trust, and social capital.

TABLE 6. Interaction between the Skill Determinants of Attitudes toward Immigration

Dep. Var: IM1 Estimation	(1) IV	(2) IV	(3) IV	(4) IV	(5) IV
yearsedu	1.066*** [0.0570]	0.943*** [0.0911]	0.228 [0.334]	0.0962 [0.326]	0.174 [0.208]
SHK		0.224* [0.116]	0.209* [0.115]	0.203* [0.116]	0.233** [0.116]
Comm.			21.88** [9.631]	22.14** [9.506]	
Manual					-38.86*** [11.17]
Observations	20098	16369	16029	15897	16029
R-squared	0.132	0.108	0.107	0.13	0.068
Attitudes	no	no	no	yes	yes

Notes:

- Standard errors clustered by region and robust to heteroskedasticity.
- Standard errors in square brackets. *** p<0.01, ** p<0.05, * p<0.1.
- Dependent variable is IM1, the answer to the question “Do you think there should be no more immigration, a few more immigrants, some more immigrants, or many more immigrants?” Question refers to immigrants belonging to the current majority group.
- All specifications include country-specific fixed effects, dummies for children, rural, and foreign-born mother. Also controls for gender, age, and age squared.
- Measures of Attitudes: Ideology, Religiosity, Happiness, Trust, and social capital.

manual skills on pro-immigration attitudes is negative and highly significant, indicating that workers employed in occupations that require high levels of manual dexterity—and hence where competition with immigrants is greater—report relatively more negative attitudes toward immigration. The effect of education remains low (and insignificant) and the point estimate for SHK increases slightly (0.23), strengthening the rejection of the null hypothesis of a zero coefficient (p-value of 0.044).

In conclusion, the two measures of protection toward immigration that we have analyzed (job-specific human capital and communication skills) are both found to have significant and distinct effects on individual attitudes toward immigration. We also find that the latter measure absorbs all the effect of education on attitudes. This is perhaps not too surprising given that the degree of communicational intensity of respondents' tasks correlates highly with years of schooling (0.48).¹⁶ In the next section we test the extent to which communicational intensity effects operate at different levels of schooling.

ROBUSTNESS

We now conduct some sensitivity analysis on our main results. Specifically, we make use of two alternative definitions of attitudes toward immigration, we re-estimate our main models on subsamples that differ in the level of education, and we experiment with substituting the measure of communication skills by the measure of manual skills.

Table 7 reports our findings. To provide a benchmark for comparison, column 1 displays the estimates from specification 3 in Table 6. Column 2 uses manual skills as a measure of comparative advantage with respect to immigrants. As noted previously, the education coefficient is drastically reduced when the regression includes either manual or communication skills. In contrast, the coefficient of job-specific skills remains largely unaltered.

Columns 3 and 4 use two alternative measures of the respondent's views toward immigration. Dependent variable IM2 is the simple average between the three questions available in the European Social Survey regarding the desired level of immigration. These questions differ on the geographic origin of the immigrants.¹⁷ The point estimate on SHK falls a bit but remains significant at 10%. Column 4 uses IM3 as dependent variable. This variable is again an average of three questions on immigration. However, these questions differ from those used in IM1 and IM2 in one key aspect. The questions used in IM3 are much broader, and report on the respondents' views on the effects of immigration on their country's economy, culture, and life in general.¹⁸ Unlike the questions used in our main analysis, these questions need not reflect the respondents' individual characteristics. That is to say, a college-educated worker and a high-school dropout may very well agree on the effects of immigration on the economy as a whole or on the country's cultural landscape. However, attending to their economic self-interest, they may sharply disagree on what immigration policy to vote for. In this light it is not surprising that the estimates in column 4 show that specific human capital does not have a significant effect on IM3. Instead years of education and manual skills appear to be significant.

The literature on comparative advantage between natives and immigrant groups focuses on low-education individuals (Peri and Sparber 2009, Amuedo-Dorantes and De la Rica 2009). In order to better compare our results to those studies, whilst at the same time deflecting the problem of high collinearity between communicational skills and education, we report estimates of our main model on two subsamples. The first subsample (column 5) contains only individuals with less than 12 years of education (57% of the population). The second subsample (column 6) includes only individuals with 12 or more years of education (43% of the total sample). Remarkably, the point estimates that we

¹⁶ The correlation between manual dexterity and years of schooling in our analytical sample is -0.31.

¹⁷ For more details, please see questions A1-A3 in Appendix 2.

¹⁸ See questions B1-B3 in Appendix 2.

TABLE 7. Robustness

Column	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Estimation	IV	IV	IV	IV	IV	IV	IV	IV
Dep. var.	IM1	IM1	IM2	IM3	IM1	IM1	IM1	IM1
yearsedu	0.228 [0.334]	0.298 [0.211]	0.359** [0.166]	0.548*** [0.173]	-0.074 [0.249]	0.254 [0.468]	0.000731 [0.175]	0.248 [0.342]
SHK	0.209* [0.115]	0.239** [0.114]	0.183* [0.0971]	0.077 [0.0932]	0.20 [0.153]	0.261* [0.138]	0.245 [0.157]	0.266* [0.139]
Comm.	21.88** [9.631]				18.53* [10.86]	25.57** [13.06]		
Manual		-38.89*** [11.24]	-43.36*** [8.916]	-44.02*** [9.581]			-30.47** [12.15]	-50.65*** [17.23]
Obs.	16029	16029	15801	15331	7858	8171	7858	8171
R-squared	0.107	0.068	0.087	0.102	0.103	0.04	0.081	0.04

Notes:

(1) Baseline. IV in our preferred specification.

(2) Manual index replaces communication intensity index.

(3) Dependent variable is average of the three questions regarding the level of immigration preferred by the respondent (IM2). See appendix.

(4) Dependent variable is average of the three questions regarding the general view on immigration (IM3). See appendix.

(5) Baseline on subsample of individuals with less than 12 years of education (57% sample).

(6) Baseline on subsample of individuals with more than 12 years of education (43% sample).

(7) Manual index replaces communicational intensity index on subsample of individuals with less than 12 years of education.

(8) Manual index replaces communicational intensity index on subsample of individuals with more than 12 years of education.

- All estimates in the table are IV. Standard errors clustered by region and robust to heteroskedasticity. Standard errors in square brackets. *** p<0.01, ** p<0.05, * p<0.1.

- All specifications include country-specific fixed effects, dummies for children, rural, and foreign-born mother. Also controls for gender, age, and age squared.

obtain are very similar in both cases and, in turn, very similar to our main estimates (columns 1 and 2). Skill-intensity does indeed seem to have attitudinal effects that are clearly distinct from those of education.

Overall, the results in this section show that that our main results are robust to the particular definition of comparative advantage used as well as to the heterogeneity in the effects of education.

CONCLUSION

This paper has re-examined the role of labor-market competition in the determination of individual preferences over immigration. A great deal of literature

has addressed this question, but has taken a very narrow view of labor market exposure to competition from immigrants. Our analysis has proposed a more comprehensive approach. In particular, we have constructed two new measures of exposure based on job-specific human capital investments and on the importance of communication-intensive tasks. Moreover, we have accounted for unobserved individual heterogeneity in tastes for immigration and for the resulting potential endogeneity in job search.

Overall our results provide strong support for the exposure hypothesis: individuals employed in jobs that are less

exposed to competition from immigrants are relatively more pro-immigration. This conclusion is supported by several findings.

First, we build a new instrument for individual exposure to labor market competition and show that it is highly relevant. The instrument is based on the assumption that the types of jobs available in one's regional labor market affect workers' actual job characteristics in ways that are unrelated to individual attitudes towards immigration. Second, our instrumental-variables estimates show that being employed in a job requiring a large amount of specific human capital leads to relatively more pro-immigration attitudes. Likewise, being employed in a communication-intensive occupation also leads to more positive attitudes toward immigration. This latter finding holds both at low and high levels of education. The observed impact of communication intensity on attitudes pro-immigration provides indirect support for the endogenous job specialization theory postulated by Peri and Sparber (2009). According to these authors, native workers respond to immigration by moving to occupations where they have a comparative advantage by virtue of being relatively better at communication-intensive tasks. Our results suggest that those native workers that manage to shift to the less exposed occupations display relatively more pro-immigration views than those that are unable to do so. In our view, supply-side constraints, such as the regional availability of low-exposure jobs play an important role in determining individual attitudes toward immigration. In conclusion, our findings suggest a larger role for labor-market competition as a determinant of attitudes toward immigration than previously found in the literature.

APPENDIX

APPENDIX 1. Observations by Country

country	Freq.	Percent	Cum.
AT	1,751	5.16	5.16
BE	1,305	3.85	9.01
CH	1,378	4.06	13.07
CZ	2,123	6.26	19.33
DE	2,026	5.97	25.30
DK	1,135	3.35	28.65
EE	1,195	3.52	32.17
ES	1,147	3.38	35.55
FI	1,494	4.40	39.96
FR	1,297	3.82	43.78
GB	1,321	3.89	47.67
GR	1,535	4.53	52.20
HU	1,156	3.41	55.61
IE	1,686	4.97	60.58
IS	465	1.37	61.95
LU	861	2.54	64.49
NL	1,364	4.02	68.51
NO	1,376	4.06	72.56
PL	1,404	4.14	76.70
PT	1,360	4.01	80.71
SE	1,431	4.22	84.93
SI	1,055	3.11	88.04
SK	1,174	3.46	91.50
TR	1,497	4.41	95.91
UA	1,386	4.09	100.00
-----+-----			
Total	33,922	100.00	

Sample: Citizens age 18-64.

APPENDIX 2. Definition Dependent Variable

We consider three variations of our dependent variable, the respondent's views toward immigration. In all cases higher values are associated to a more pro-immigration stance.

Our main dependent variable is IM1. It is based on the following question:

B35~ CARD 14 Now, using this card, to what extent do you think [country] should¹⁹ allow people of the same race or ethnic group as most [country's] people to come and live here²⁰?

Allow many to come and live here	1
Allow some	2
Allow a few	3
Allow none	4
(Don't know)	8

We re-scale the values to range between 25 (allow none) to 100 (allow many).

The second variation of our dependent variable is IM2. This variable is a simple average of questions B35, B36 and B37. It also ranges from 25 to 100.

B36~ STILL CARD 14 How about people of a different race or ethnic group from most [country] people? Still use this card.

B37~ STILL CARD 14 How about people of a poorer countries outside Europe? Use the same card.

The third variation (IM3) is based on three questions regarding the respondent's general view on immigration. Unlike the previous set of questions, questions B38-B40 are what political scientists call *sociotropic*, that is, they refer to the effects of immigration for the country as a whole and hence we believe they are less clearly related to the effects of immigration on the respondent's individual labor market outcomes. As before, we take a simple average and re-scale the variable to take on a maximum value of 100.

B38~ CARD 15 Would you say it is generally bad or good for [country]'s economy that people come to live here from other countries? Please use this card.

Bad for the economy										Good for the economy	(Don't Know)
00	01	02	03	04	05	06	07	08	09	10	88

¹⁹ "Should" in the sense of 'ought to'; not in the sense of 'must'.

²⁰ "Here" = country throughout these questions.

B39~ CARD 16 And, using this card, would you say that [country]'s cultural life is generally undermined or enriched by people coming to live here from other countries?

Cultural life is undermined									Cultural life is enriched	(Don't Know)	
00	01	02	03	04	05	06	07	08	09	10	88

B40~ CARD 17 Is [country] made a worse or a better place to live by people coming to live here from other countries? Please use this card.

Worse place to live									Better place to live	(Don't Know)	
00	01	02	03	04	05	06	07	08	09	10	88

APPENDIX 3. Tabulation Main Variables

APP3.1. Dependent Variable (IM1)

Original question in ESS2004:

allow many/few immigrants of same race/ethnic group as majority	Freq.	Percent	Cum.
allow many to come and live here	4,476	22.25	22.25
allow some	9,239	45.92	68.16
allow a few	5,015	24.92	93.09
allow none	1,391	6.91	100.00
Total	20,121	100.00	

Our transformation:

imdv1	Freq.	Percent	Cum.
25	1,391	6.91	6.91
50	5,015	24.92	31.84
75	9,239	45.92	77.75
100	4,476	22.25	100.00
Total	20,121	100.00	

APP3.2. Job-Specific Human Capital

Original question in ESS2004

somebody with right qualification, how long to learn to do your job well	Freq.	Percent	Cum.
1 day or less	669	3.99	3.99
2-6 days	1,598	9.54	13.54
1-4 weeks	2,819	16.83	30.37
1-3 months	3,772	22.52	52.89
more than 3 months, up to 1 year	4,517	26.97	79.87
more than 1 year, up to 2 years	1,993	11.90	91.77
more than 2 years, up to 5 years	1,063	6.35	98.11
more than 5 years	316	1.89	100.00
Total	16,747	100.00	

Our transformation (in months)

shk_months	Freq.	Percent	Cum.
.0166667	669	3.99	3.99
.1333333	1,598	9.54	13.54
.5833333	2,819	16.83	30.37
2	3,772	22.52	52.89
7.5	4,517	26.97	79.87
18.25	1,993	11.90	91.77
42.58333	1,063	6.35	98.11
73	316	1.89	100.00
Total	16,747	100.0	

APP3.3. Communication Intensity Index

Communicational Skills by occupation			
(O*NET)	Freq.	Percent	Cum.
.09	464	2.31	2.31
.108	454	2.26	4.56
.12	128	0.64	5.20
.123	81	0.40	5.60
.148	538	2.67	8.27
.188	175	0.87	9.14
.207	1,013	5.03	14.18
.212	1,037	5.15	19.33
.242	704	3.50	22.83
.247	819	4.07	26.90
.26	916	4.55	31.45
.363	1,049	5.21	36.66
.469	1,692	8.41	45.07
.482	1,630	8.10	53.17
.52	880	4.37	57.55
.538	474	2.36	59.90
.56	283	1.41	61.31
.638	685	3.40	64.71
.657	1,768	8.79	73.50
.687	691	3.43	76.93
.711	878	4.36	81.30
.715	1,115	5.54	86.84
.759	1,086	5.40	92.23
.76	443	2.20	94.43
.807	1,067	5.30	99.74
.95	53	0.26	100.00
Total	20,123	100.00	

APP3.4. Manual Dexterity Index

Manual skills by occupation			
(O*NET)	Freq.	Percent	Cum.
.244	878	4.36	4.36
.276	53	0.26	4.63
.328	1,086	5.40	10.02
.334	1,115	5.54	15.56
.344	1,768	8.79	24.35
.359	283	1.41	25.76
.379	474	2.36	28.11
.39	1,067	5.30	33.41
.426	1,630	8.10	41.51
.465	916	4.55	46.07
.487	691	3.43	49.50
.509	1,692	8.41	57.91
.517	1,049	5.21	63.12
.536	443	2.20	65.32
.623	464	2.31	67.63
.631	685	3.40	71.03
.679	704	3.50	74.53
.693	1,699	8.44	82.97
.708	81	0.40	83.38
.741	1,013	5.03	88.41
.756	175	0.87	89.28
.793	538	2.67	91.95
.83	1,037	5.15	97.11
.841	454	2.26	99.36
.86	128	0.64	100.00
Total	20,123	100.00	

APPENDIX 4. Construction Indices for Communication and Manual Intensities

O*NET reports descriptors for up to 965 different occupations when occupations are coded at the maximum level of detail (i.e. 4-digits). Yet US occupations are much harder to match into the International System of Occupational Coding, ISCO-88, when they are coded in 4-digits. Hence we use an O*NET dataset where occupations are coded at 3 digits (N= 449). Moreover, crosswalks from the occupational coding system deployed by O*NET and ISCO-88 have so far only been available for a 3-digit into 2-digit conversion. ISCO-88 at 2-digits groups the 449 O*NET occupations into 26 broader occupational categories.

For the construction of our indices we proceed as follows: First, informed by factor analysis, we compute communicational intensity and manual intensity scores using the 3-digit O*NET dataset. This procedure, which is described in detail below, assigns both a communicational intensity and a manual intensity score to each of the 449 occupations of the O*NET dataset. Secondly, using the crosswalk kindly provided by Sara De la Rica and Catalina Amuedo-Dorantes, we convert our O*NET occupations into their ISCO-88 2-digit equivalents. All descriptors, as well as our two measures of communication and manual skill intensity are converted into 2-digits by averaging the scores of all 3-digit occupations belonging to the same 2-digit category. Finally, we match this information to the ESS data. Our final skill measures are thus much less refined than it would be in principle possible, should we have a 3-to-3-digit converter (or even a 4-to-4). Information is undoubtedly loss in the matching process. For example, O*NET provides detailed skill descriptors for “Elevator installers and repairers”, yet in order to match this US 3-digit occupation into the ESS, we need to average these skill descriptors with those of all other occupations belonging to the same ISCO-88 2-digit equivalent, which is “Metal, machinery and related trade workers”. Loss of information is unavoidable giving our matching procedure.

The construction of the skill indices in O*NET was informed by exploratory factor analysis. Principal-component factor analysis with orthogonal varimax rotation identified 10 different factors out of the 277 skill descriptors in the O*NET dataset (at 3-digits). The first factor accounted for 25.6 of the variance. Skill-descriptors with rotated factor loadings higher than 0.6 in this first factor included both skills directly involved in communication (i.e. oral comprehension, oral expression, written comprehension, written expression, speech recognition, and speech clarity) as well as skills relating to abstract thinking (i.e. fluency of ideas, originality, problem sensitivity, deductive and inductive reasoning, information ordering, category flexibility, memorization, etc.). Our communication intensity measure only uses the 6 former skills, since there is no reason to suppose any comparative advantage of natives in the latter. The index is the result of averaging the task-importance and the observed-ability scores of these 6 communicational skills for each of the 449 occupations²¹ —which were later condensed into their ISCO-88 2-digit equivalents. Although we only focus on direct communicational skills, it must be noted that occupations where communicational skills are on demand tend to be those that also require abstract thinking.

The fourth factor of the principal component analysis identified 7 different skills/abilities relating to physical dexterity (i.e. visualization, arm-hand steadiness, manual dexterity, finger dexterity, control precision, wrist-finger speed and visual color discrimination). These 7 skills correspond to the descriptors with rotated factor scores higher than 0.55. The manual dexterity factor accounted for 10.1 per cent of the skill variance. The manual intensity index is constructed by averaging the task-importance and observed ability scores of these 7 skills.

²¹ For each skill involved in any given occupation, O*NET experts evaluate 2 different dimensions: 1) how important is this given skill/ability for the occupation and 2) the average observed levels of such skill in the occupation. Both dimensions correlate very highly, as can be expected. Factor analysis and hence our indices use both type of descriptors so for each of the skills involved in our indices we actually average 2 different descriptors, one referring to task-importance and the other referring to observed levels.

APP4.1. The Top 10 O*NET Occupations (3 Digits) with Highest Communicational-Intensity and Manual Intensity Scores

Communicational skills	Score	Manual skills	Score
Public Relations Specialists	0.96	Electricians	0.97
Chief executives and legislators	0.95	Cabinetmakers and Bench Carpenters	0.97
Clergy	0.95	Jewelers and Precious Stone Workers	0.97
Human Resources Managers	0.93	Structural Iron and Steel Workers	0.96
Social and Community Service Managers	0.93	Engine and Other Machine Assemblers	0.95
Speech-Language Pathologists	0.93	Electronic Equip. Installers & Repairers Motor Vehicles	0.95
Advertising and Promotions Managers	0.92	Packaging and Filling Machine Operators and Tenders	0.95
Paralegals and Legal Assistants	0.92	Electronic Home Entertainment Equipment Installers and Repairers	0.95
Podiatrists	0.92	Electric Motor, Power Tool, and Related Repairers	0.94
Sales Engineers	0.92	Tool Grinders, Filers, and Sharpeners	0.94

APPENDIX 5. Ordered Probit Estimates for Main Models

Dep. Var: IM1 Estimation	(1) IV	(2) IV	(3) IV	(4) IV	(5) IV
yearsedu	0.087*** [0.009]	0.008 [0.306]	0.015 [0.019]	-0.03 [0.035]	0.008 [0.042]
SHK	0.024** [0.010]	0.024** [0.001]	0.028** [0.095]	0.024** [0.012]	0.026** [0.012]
Comm.		2.514*** [0.946]		1.77* [1.066]	3.26*** [1.16]
Manual			-4.478*** [1.046]		
Observations	20098	20098	16029	10146	9952
Pseudo R-squared	0.0601	0.0605	0.0611	0.048	0.051

Notes:

- (1) Years of schooling and specific human capital.
- (2) Years of schooling, specific human capital and communication intensity. Baseline.
- (3) Manual index replaces communication intensity index.
- (4) Baseline on subsample of individuals with less than 12 years of education (57% sample).
- (5) Baseline on subsample of individuals with more than 12 years of education (43% sample).

- All estimates in the table are IV. Standard errors clustered by region and robust to heteroskedasticity. Standard errors in square brackets. *** p<0.01, ** p<0.05, * p<0.1.
- All specifications include country-specific fixed effects, dummies for children, rural, and foreign-born mother. Also controls for gender, age, and age squared.

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